

Moment Generating Functions

Definition and properties of MGF

For a probability density $f[x] \geq 0$ with $\int_{-\infty}^{+\infty} f[x] dx = 1$ we define the MGF (moment generating function)

$$g[r] = \int_{-\infty}^{+\infty} e^{rx} f[x] dx$$

which exists uniquely for each real number r since the integrand is not negative. The integral may however take the value $+\infty$ (i.e. the integral defining the moment generating function may for some r be divergent).

The same considerations apply to a discrete density $p[x] \geq 0$ with $p[x] > 0$ only for enumerably many values of x and $\sum_x p[x] = 1$.

If the integral $g[r]$ is finite for some value $r_1 < 0$ and also for some other value $r_2 > 0$ then it is known that $g[r]$ is necessarily finite for every r between these values. It implies that $g[r]$ will necessarily be finite on an interval neighborhood of the origin $r = 0$. In such a case all derivatives of g exist at $r = 0$ and may be computed under the integral! This means that in such a case (for every k) the k -th derivative at $r = 0$ exists and satisfies

$$\frac{d^k}{dr^k} g[r] \Big|_{r=0} = \int_{-\infty}^{+\infty} \frac{d^k}{dr^k} e^{rx} f[x] dx$$

$$= \int_{-\infty}^{+\infty} x^k e^{0x} f[x] dx = E X^k .$$

That is, we can obtain the integer moments of r.v. X !!!

Example: MGF for Bernoulli

Bernoulli r.v. X with $P(X = 1) = p$, $P(X = 0) = q$. We have discrete probability density $p[0] = q$, $p[1] = p$ and MGF

$$g[r] = p e^{r1} + q e^{r0} = p e^r$$

which is finite for every real number r . Hence g is finite for values of r in a neighborhood of 0.

Let us find $E X$ from the MGF. We have $E X = g'[0] = p e^0 = p$.

Let us find $E X^2$ from the MGF. It is $E X^2 = g''[0] = p e^0 = p$.

Therefore the variance $\text{Var } X = E X^2 - (E X)^2 = p - p^2 = p q$.

Example: MGF for Binomial

Binomial distribution n, p . We have r.v. $S = X_1 + \dots + X_n$ where r.v. X_i are i.i.d. Bernoulli p . Because of independence we have

$$\begin{aligned} E e^{rS} &= E e^{r(X_1 + \dots + X_n)} = E e^{rX_1} \dots E e^{rX_n} \\ &= g[r] \dots g[r] = (p e^r + q)^n \end{aligned}$$

Let us find $E S$ (which we know to be np) from the above MGF for S which is (different from Bernoulli) $g[r] = (p e^r + q)^n$.

$$\begin{aligned} E S &= \frac{d}{dr} (p e^r + q)^n \text{ (at } r = 0) \\ &= n (p e^r + q)^{n-1} p e^r \text{ (at } r = 0) \\ &= n (p e^0 + q)^{n-1} p e^0 = np. \end{aligned}$$

Let us find $E S^2$ from the MGF for r.v. S .

$$\begin{aligned} E S^2 &= \frac{d^2}{dr^2} (p e^r + q)^n \text{ (at } r = 0) \\ &= \frac{d}{dr} n (p e^r + q)^{n-1} p e^r \text{ (at } r = 0) \\ &= n p [(n-1) (p e^r + q)^{n-2} p e^r + (p e^r + q)^{n-1} e^r] \end{aligned}$$

at $r = 0$ is

$$= n p [(n-1) p + 1]$$

So $\text{Var } S$ (which we know to be npq) is by the MGF

$$\begin{aligned} \text{Var } S &= E S^2 - (E S)^2 = n p [(n-1) p + 1] - (np)^2 \\ &= n p [-p + 1] = n p q. \end{aligned}$$

Using *Mathematica* to evaluate the derivatives.

```
In[1]:= g[x_] := (p Exp[r] + q) ^ n
```

```
In[2]:= ? D
```

D[f, x] gives the partial derivative of f with respect to x. D[f, {x, n}] gives the nth partial derivative of f with respect to x. D[f, x1, x2, ...] gives a mixed derivative.

```
In[3]:= D[g[r], {r, 1}]
```

```
Out[3]= E^r n p (E^r p + q)^(-1+n)
```

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In[4]:= D[g[r], {r, 2}]
```

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Out[4]= E^(2 r) (-1 + n) n p^2 (E^r p + q)^(-2+n) + E^r n p (E^r p + q)^(-1+n)
```

```
In[5]:= ? Series
```

Series[f, {x, x0, n}] generates a power series expansion for f about the point x = x0 to order (x - x0)^n. Series[f, {x, x0, nx}, {y, y0, ny}] successively finds series expansions with respect to y, then x.

Series[f, {x, x0, n}] generates a power series expansion for f about the point x = x0 to order (x - x0)^n.

```
In[6]:= Series[g[r], {r, 0, 2}]
```

```
Out[6]= (p + q)^n + n p (p + q)^(-1+n) r + (1/2 (-1 + n) n p^2 (p + q)^(-2+n) + 1/2 n p (p + q)^(-1+n)) r^2 + O[r]^3
```

So we see from the series that the coefficient of "r" is

$$n p (p + q)^{n-1} = n p = E S \quad (\text{since } p + q = 1)$$

and also TWICE the coefficient of " r^2 " is

$$(n-1) n p^2 + n p = n p ((n-1) p + 1)$$

which we got above. So Var S = npq by this Series method also.

MGF for the uniform on interval [a, b]

The MGF for the uniform on [a, b] is

$$g[r] = \frac{1}{b-a} \left(\frac{e^{br}}{r} - \frac{e^{ar}}{r} \right)$$

We expect the derivatives of g to have indeterminate forms $0/0$ which may confuse *Mathematica*.

```
In[7]:= g[r_] :=  
      (Exp[b r] - Exp[a r]) / (r (b - a))
```

```
In[8]:= D[g[r], {r, 1}]
```

$$\text{Out[8]= } -\frac{-E^{a r} + E^{b r}}{(-a + b) r^2} + \frac{-a E^{a r} + b E^{b r}}{(-a + b) r}$$

```
In[9]:= D[g[r], {r, 2}]
```

$$\text{Out[9]= } \frac{2 (-E^{a r} + E^{b r})}{(-a + b) r^3} - \frac{2 (-a E^{a r} + b E^{b r})}{(-a + b) r^2} + \frac{-a^2 E^{a r} + b^2 E^{b r}}{(-a + b) r}$$

So the derivatives need careful examination to see what happens at $r = 0$. We resort to having *Mathematica* carry out a Taylor's Series expansion approximation of $g[r]$ near $r = 0$.

```
In[10]:= Series[g[r], {r, 0, 2}]
```

$$\text{Out[10]= } 1 + \frac{\left(-\frac{a^2}{2} + \frac{b^2}{2}\right) r}{-a + b} + \frac{\left(-\frac{a^3}{6} + \frac{b^3}{6}\right) r^2}{-a + b} + O[r]^3$$

Now we can "read off" the mean and second moment as the series coefficient of "r" and twice that of "r²" respectively. So

$$E U = \frac{\left(-\frac{a^2}{2} + \frac{b^2}{2}\right)}{-a+b} = \frac{a+b}{2}$$

$$E U^2 = 2 \frac{\left(-\frac{a^3}{6} + \frac{b^3}{6}\right)}{-a+b} = (a^2 + a b + b^2) / 3$$

From these we find that $\text{Var } U = (b - a)^2 / 12$.

Some Example Questions

1. For the standard normal distribution the MGF is $g[r] = e^{\frac{r^2}{2}}$ for every real number r . Find the mean, second moment, and variance.

Here is a solution using Mathematica to do the derivatives.

```
In[11]:= g[r_] := Exp[r^2 / 2]
```

```
In[12]:= D[g[r], {r, 1}]
```

```
Out[12]= E^{\frac{r^2}{2}} r
```

```
In[13]:= D[g[r], {r, 2}]
```

```
Out[13]= E^{\frac{r^2}{2}} + E^{\frac{r^2}{2}} r^2
```

From these derivatives we can evaluate at $r = 0$ to obtain

$$E Z = g'[0] = e^0 \cdot 0 = 0$$

$$E Z^2 = g''[0] = e^0 + e^0 \cdot 0 = 1$$

so $\text{Var } Z = 1 - 0^2 = 1$.

Using the Taylor's Series expansion for $g[r]$ around $r = 0$ we can "read off" many moments.

```
In[17]:= Series[g[r], {r, 0, 12}]
```

$$\text{out[17]} = 1 + \frac{r^2}{2} + \frac{r^4}{8} + \frac{r^6}{48} + \frac{r^8}{384} + \frac{r^{10}}{3840} + \frac{r^{12}}{46080} + O[r]^{13}$$

All of the odd moments are zero since the coefficients of odd powers of r are all zero. This is no surprise since the standard normal distribution is symmetric around $r = 0$. But we read off

$$E Z^4 = 4! (\text{coeff of } r^4) = 4! / 8 = 3$$

$$E Z^6 = 6! (\text{coeff of } r^6) = 6! / 48 = 15$$

(there is a simple pattern to these).

2. The MGF for the Poisson- λ distribution is

$$g[r] = e^{\lambda(e^r - 1)}$$

for every real number r . We find that

$$g'[r] = e^{\lambda(e^r - 1)} \lambda e^r$$

so the mean of the Poisson is $g'[0] = \lambda$.

The second moment is gotten from

$$g''[r] = e^{\lambda(e^r-1)} \lambda e^r + e^{\lambda(e^r-1)} (\lambda e^r)^2$$

from which we find that the second moment is $g''[0] = \lambda + \lambda^2$. Thus we find (for the Poisson) the variance to be the same as the mean λ .

Expanding $g[r]$ in series we get

```
In[19]:= g[r_] := Exp[λ (Exp[r] - 1)]
```

```
In[21]:= Series[g[r], {r, 0, 2}]
```

```
Out[21]= 1 + λ r + 1/2 (λ + λ^2) r^2 + O[r]^3
```

We read off the mean λ (the coeff of r) and the second moment ($\lambda + \lambda^2$) (twice the coeff of r^2). So the variance of the Poisson is also equal to the mean λ .

3. The MGF of the exponential (waiting time) distribution with parameter λ is

$$g[r] = \frac{\lambda}{\lambda - r}$$

for every real number $r < \lambda$. Since $\lambda > 0$ this means that we have all moments and

$$E T = g'[0] = \frac{1}{\lambda}$$

$$E T^2 = g''[0] = \frac{2}{\lambda^2}$$

So $\text{Var } T = \frac{1}{\lambda^2}$. We can also consult the series to look at the moments.

$$\mathbf{g[r_]} := \lambda / (\lambda - \mathbf{r})$$

```
In[22]:= Series[g[r], {r, 0, 2}]
```

$$\textit{out[22]} = 1 + \lambda r + \frac{1}{2} (\lambda + \lambda^2) r^2 + O[r]^3$$

Notice that the computer program has not informed us that $g[r]$ is infinite for $r > \lambda$.